

# Information Technology and Political Engagement: Mixed Evidence from Uganda

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## **Abstract**

This study integrates three related field experiments to learn about how Information Communications Technology (ICT) innovations can affect who communicate with politicians. We implemented a nationwide experiment in Uganda following a smaller-scale framed field experiment that showed that ICTs can lead to significant flattening: a greater share of marginalized populations used SMS-based communication compared to existing political communication channels. We find no evidence for such flattening from the national experiment. Instead, patterns of participation look like politics as usual: participation rates are low and marginalized populations engage at especially low rates. We examine possible reasons for these differences, and then present the design and analysis of a third mechanism experiment that helps parse rival explanations for these divergent patterns. The evidence suggests that even when citizens have issues they want to raise, technological fixes to communication deficits can be easily undercut by structural weaknesses in political systems.

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# 1 Introduction

Weak political communication channels characterize many developing countries. Traditional aggregators of interests, such as political parties, have limited reach, and regular public opinion polls are all but non-existent. Many citizens have only limited opportunities to directly communicate with politicians, usually around elections, and only very few are willing to bear the high costs of reaching out to representatives to articulate interests outside electoral cycles. In turn, weak political communication channels have important implications for the health of a country’s democratic institutions: with poor information on their constituents’ preferences and policy priorities, elected representatives have a hard time representing interests, and political parties cannot differentiate themselves in meaningful ways. The starting point of this study is that strengthening weak political communication channels offers a promising way to begin improving political representation.

In this paper, we report findings from a multi-year research project (involving three related field experiments) designed to test whether innovations in information communication technologies (ICTs) can be harnessed to improve political communication in low-income countries. Since the existence and costs of new ICT platforms are likely correlated with features of a political system that may independently determine political engagement, assessing the effects of technological innovations on political communication is fraught with difficulties. To overcome this identification challenge, we partnered with the Parliament of Uganda and the National Democratic Institute (NDI), an international non-government organization (NGO), to implement one of the largest field experiments involving political elites to date.

The primary experiment examines a nationwide Parliament-led program that introduced a new channel for contacting elected representatives. In the terminology of Harrison and List (2004), this experiment is a “natural field experiment” (NFE), implemented as part of the political process in Uganda. The intervention established and subsidized a mobile-technology platform for political communication with the goal of increasing and diversifying citizen voice. Citizens in over 100 treatment constituencies were able to communicate with their Member of Parliament (MP) by sending text-messages at low (or no) cost.

MPs representing treated constituencies could respond to messages via the platform and use the system’s functionalities to aggregate messages and to learn about usage patterns over time. The ICT platform was introduced to voters via twice daily short radio ads in nineteen national languages over a six-months period. This experiment is unusual in scale—the program involved about 10 million voters—but also in nature: change in access was led by political elites and thus provided a relatively strong invitation to citizens to engage in politics.

The results of the nation-wide field experiment are disappointing: uptake in treatment constituencies was low, and marginalized populations largely refrained from using the ICT platform. In fact, because of the disappointing level of citizen engagement and revealed low interest among Members of Parliament (MPs), the Ugandan Parliament ultimately decided to phase out the SMS service.

Importantly, these disappointing findings differed markedly from findings from a more controlled experiment—in the terms of Harrison and List (2004), a “framed field experiment” (FFE)—undertaken before the national program was rolled out. Results from the FFE suggested a relatively high demand, and that mobile technology could democratize political communication because marginalized constituents were willing to engage at relatively high rates and were not more price sensitive, compared to less marginalized voters. By contrast, the NFE found little citizen involvement and no improvement in *differential* access to political elites.

In the second part of the paper we take advantage of differences between the NFE and the FFE to explore the reasons for the disappointing findings in the national experiment. Since both experiments were implemented using subjects from constituencies across Uganda, they involved similar populations, eliminating common external validity concern—that replications tend to fail because of unobserved features of the experimental subject pool (Allcott 2015). Instead, we find relatively strong evidence that voters doubt the efficacy of contacting their MP directly, and suggestive evidence that larger (structural) inequalities prevented the ICT program from having effects at scale. We also find evidence suggesting that the tools governments have at their disposal for informing citizens matters, with different marketing strategies employed over these experiments inducing varying invitational effects (and thus, variation in take-up conditional on hearing about the service). Specifically, our findings cast doubt on the utility of using short radio ads to elicit wide-scale participation. We find no evidence, however, that scale itself is driving our core results.<sup>1</sup>

This paper makes several contributions to the literature on political communication, and especially to our understanding of inequalities in political participation. We highlight how the underlying willingness to engage in politics—even when using low-cost impersonal communication channels—crucially depends on citizen beliefs about the effectiveness of political engagement, which itself likely depends on politicians’ response to incoming messaging.<sup>2</sup> Though not identified, we provide below evidence that the usage of the system was tightly connected to MP’s (in)action. ICTs, we argue, in and of themselves, do not make non-responsive politicians responsive.

The paper also contributes to a growing literature on the effectiveness of using ICT innovations to improve governance outcomes. Past studies have focused on using ICTs to reduce absenteeism among frontline service providers (Duflo, Hanna, and Rya 2012; Grossman, Platas, and Rodden forthcoming), improve election integrity (Callen and Long 2014), increase engagement in local affairs (Buntaine, Daniels, and Devlin 2018), and report corruption (Blair, Littman, and Paluck 2017) and violent incidents (van der Windt, Peter and Humphreys, Macartan 2016). Ours is the first study to examine the role ICTs may play in altering citizen-MPs relationships in the context of low-income countries.

Our study also contributes to ongoing methodological debates on the utility of relatively small-scale controlled experiments, such as the framed field experiment described here (and, a fortiori, “artefactual” field experiments or lab experiments), in shedding light on core political processes. Most field experiments—including many natural field experiments—are implemented on a small scale but seek to make claims about large-scale processes. For example, small-scale experiments may be used to test new approaches, be designed as a proof of concept, or test micro-logics that arguably underlie general features of human behavior. Indeed, much of the “credibility revolution” in the study of international development is premised on the idea that small-scale field experiments can create a body of knowledge that allows promoting “what works” and eliminating programs and policies that do not (Banerjee and Duflo 2009). Yet, it is often contestable whether the results of small-scale field experiments can accurately inform theory or form the basis for more general policy (Manski 2013). Our study distinguishes between explanations for when and why such inferences may not be valid and garners evidence for or against these different explanations.

In the remainder of this paper we introduce the research questions that the different field experiments were designed to answer and present the design and results from the scaled-up national program. We then present

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<sup>1</sup>Plausibly, radio programming may be effective even if radio ads are not (Yanagizawa-Drott 2014; Adena et al. 2015).

<sup>2</sup>For recent studies making similar claims, see Sjoberg, Mellon, and Peixoto (2017) and Grossman, Michelitch, and Santamaria-Monturiol (2017).

analyses designed to assess mechanisms that could account for differences in outcomes. Our conclusions focus on the implications for efforts to democratize political communication, and on the implications for learning about political processes from controlled experiments.

## 2 Access as a Constraint on Political Communication

In many low-income countries, the aggregation of preferences is limited by the weakness of civil society organizations, labor unions and political parties. Potential preference aggregators, such as unions and non-government organizations (NGOs), tend to be located in urban centers and to have a narrow membership base. Political parties may have a wider reach, but many are weakly institutionalized and lack resources and elite cohesion (LeBas 2011). That parties are often organized on ethnic or geographic basis—rather than by class or religion—may in turn contribute to the non-programmatic nature of many political parties, in Africa and beyond (Riedl 2014).

The political implications of weak preference aggregations are manifold. Lacking information on voters’ policy priorities, quintessential preference aggregators such as political parties, often focus on valence issues that offer voters little policy differentiation (Bleck and Walle 2013). When parties are non-programmatic, the accountability relationship between office holders and voters can narrow down to local clientelistic exchange (Stokes et al. 2013).

While political parties commonly lack the capacity to elicit citizens’ preferences systematically, and on a wide-scale, constituents too may be reluctant to bear the costs of political communication. Constituents likely will not invest in articulating preferences if they doubt that government officials would be responsive to citizen demands. This sort of low sense of (external) efficacy may be especially prevalent where governments have low capacity and/or low levels of legitimacy (Craig, Niemi, and Silver 1990). A sense of (internal) efficacy—i.e., the belief that one has the personal ability to participate effectively in politics (Niemi, Craig, and Mattei 1991)—can be especially weak for marginalized populations, whether defined by gender, education, wealth or partisanship (Coleman and Davis 1976).<sup>3</sup> Notably, a weak sense of political efficacy is compounded by the high cost of traditional forms of political communication—e.g., traveling large distances to meet public officials in person—that further reduce citizens’ incentive to proactively reach out to politicians in order to articulate interests, needs and policy preferences.

### 2.1 Logic

A simple model clarifies the implications arising from political bias and information asymmetries and helps motivate some of the mechanisms we examine in Section 5. Our goal is to explore the logic of preference articulation in weak information environments, and especially whether the incentives to bear the costs of political communication might differ across subsections of the population.

Consider a politician who has to decide what share  $\alpha_i^j$  of resources to allocate to group  $i$  in sector  $j$  subject to  $\sum_i \sum_j \alpha_i^j = 1$ . Say that each group  $i$  values only one sector but the politician is uncertain about the

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<sup>3</sup>Although voting rates are sometimes higher among the poor (Kasara and Suryanarayan 2015) and less educated (Croke et al. 2016), this often reflects differences in mobilization (or repression) in different contexts and does not extend immediately to other types of political engagement.

sectoral preference of groups and believes group  $i$  favors sector  $j$  with probability  $q_i^j$ . Politicians maximize a weighted average of expected group welfare with bias parameter  $\beta_i$ ,  $\sum_i \beta_i = 1$ :

$$u(\alpha) = \sum_i \beta_i \sum_j q_i^j \sqrt{\alpha_i^j}$$

This set up can represent a large range of institutional environments. By letting one of the groups correspond to the politician themselves, or their party, we can capture variation in the degree to which politicians seek to respond to the interests of constituents. Similarly, the  $\beta_i$  term can be interpreted as capturing either the electoral importance of groups or non-electoral significance, such as ethnic affinity with politicians, which past work has demonstrated can affect distributional outcomes (Burgess et al. 2015; Kramon and Posner 2016). The square root function here may be interpreted either as reflecting concavity in the benefits to groups from a policy or in the politician’s valuation of these benefits to voters. The main results below extend immediately to more general formulations such as  $\sum_i \beta_i \sum_j q_i^j (\alpha_i^j)^\gamma$ . The formulation does not however handle decision rules in which the gains from supporting constituents depend on how satisfied other constituents are, as is the case for coalitional politics (Humphreys 2008).

The question of interest then is what preferences do politicians have over information on voters? What preferences do voters have over politicians’ information and how do these affect equality of outcomes?

Given the above utility function, the politician’s optimal allocation is:

$$\alpha_i^j = \frac{(\beta_i q_i^j)^2}{\sum_h \sum_k (\beta_h q_h^k)^2}$$

Thus distributions to groups reflect how informed politicians are about group preferences: they allocate more where they can allocate accurately. The total allocation to a group is then  $\frac{\beta_i^2 \sum_j (q_i^j)^2}{\sum_h \sum_k (\beta_h q_h^k)^2}$ , which, along with bias, depends on the *fragmentation* of beliefs—corresponding here to the squared term on  $q$  in the numerator—the more fragmented beliefs are across sectors the less is provided to a group overall.

Consider a case with two group and two sectors. If there were certainty that group 1 preferred sector 1 and uncertainty over group 2’s preferences, group 1 would receive twice as many benefits as group 2 *even in the absence of ethnic, or other sources of, bias*:  $\sum_j \alpha_i^j = \frac{2\beta_1^2}{2\beta_1^2 + \beta_2^2}$ .

In this case a politician’s “indirect” utility, given information  $q$  and under the assumption that she implements optimal allocations, would be:

$$v(q) = \sum_i \beta_i \sum_j q_i^j \left( \frac{(\beta_i q_i^j)^2}{\sum_h \sum_k (\beta_h q_h^k)^2} \right)^{.5} = \left( \sum_i \sum_k (\beta_i q_i^k)^2 \right)^{.5}$$

We see from this that, *ceteris paribus*, the marginal gains for a politician from a reduction in the fragmentation of beliefs is greater for favored groups. For example, starting from an uninformed position, if a politician were to choose between being informed about group 1 or group 2, this would mean a comparison between  $(\beta_1^2 + .5\beta_2^2)^{.5}$  and  $(\beta_2^2 + .5\beta_1^2)^{.5}$ . The former exceeds the latter if  $\beta_1 > \beta_2$ .

Thus if they have to choose, politicians invest more in learning about the preferences of favored groups.

Because of this, *favoured groups may be benefited doubly*: because politicians *care* more about them and because they *know* more about them. These simple logics highlight how in the presence of bias there can be inequality in information, which contributes to greater inequality in resource allocation.

Say now that voters can take actions to render politicians more informed. Will they have incentives to do so? The sensitivity of allocations to a voter 1's preferred sector (here across two sectors) to beliefs that 1's preferred sector is indeed sector 1 is given by:

$$\frac{\partial \alpha_1^1}{\partial q_1^1} = 2\beta_1^2 q_1^1 \frac{\sum_{k \neq 1} \sum_h (\beta_k q_k^h)^2 + (1 - q_1^1) \beta_1^2}{(\sum_k \sum_h (\beta_k q_k^h)^2)^2} \quad (1)$$

This is positive everywhere, as long as  $\beta_1 > 0$ , even if  $\beta_1 < \beta_2$ .

Thus there are always gains for a voter from politicians being more informed about their preferences—more information will always mean better targeting of resources. Note that it is quite possible that better information means that the politician is more uncertain about priorities (for example if they started putting greater weight on the wrong sector) and will allocate less to a group, but still allocate it more effectively to that group. It is also ambiguous whether the gains to more information are increasing in the bias of the politician towards other groups. Similarly, marginal gains are lower when politicians are more informed about other groups (i.e. have less dispersed beliefs).<sup>4</sup>

There can however be ranges in which groups for whom there is weak information have greater incentives to provide information than groups about whom there is good information, even if biases go against them. To see this, note that with two groups (and dropping the sector superscripts on  $q_i$ ), the marginal gains are greater for group 1 if:

$$\beta_1^4 q_1 (1 - q_1) + \beta_1^2 \beta_2^2 (q_1 (q_2^2 + (1 - q_2)^2) - q_2 (q_1^2 + (1 - q_1)^2)) - \beta_2^4 q_2 (1 - q_2) > 0$$

This has a term that is increasing in the preference weighted uncertainty of own group preferences, decreasing in the preference weighted uncertainty of other group preferences and a third term that depends on relative uncertainty and relative preferences (increasing in the fragmentation of other group information and decreasing in the fragmentation of own group information).

Thus a weaker group may have greater incentive to communicate if there is greater uncertainty about their preferences and less uncertainty about the privileged groups preferences. The point is obvious at the extremes, if  $q_2 = 1$  the condition is  $q_1 \beta_1^2 > (1 - 2q_1) \beta_2^2$  which holds for any  $q_1 \geq .5$ .

Figure 1 (upper panel) sharpens this intuition for a case with two sectors and preferences for the privileged group of  $\beta_2 = .55$  and for the marginalized group of  $\beta_1 = .45$ . Here, with prior beliefs on the privileged group's preferred sector of  $q_2 = 0.9$  the marginalized group has a stronger incentive to provide information than the privileged group in a set of situations in which the politician is (strictly) less informed about it. The lower panel illustrates the depressing effect of relative marginalization in most ranges – in general the more privileged is the privileged group the weaker is the incentive for the marginalized group to provide information.

<sup>4</sup>Note that the marginal gain has the form:  $a(x+b)(x+d)^{-2}$ , so the derivative is:  $a(x+d)^{-2} - 2a(x+b)(x+d)^{-3} = a \frac{d-2b-x}{(x+d)^3}$ , since  $x+d$  is positive, the derivative is negative as long  $d < x+2b$ . In general, this cannot be guaranteed. For example, with  $x=0$ , this condition is equivalent to  $\beta_1^2((q_1)^2 + (1-q_1)^2) < 2\beta_2^2(1-q_1)$  or  $q_1 < .5$ <sup>5</sup>. See Figure 1 for intuition.

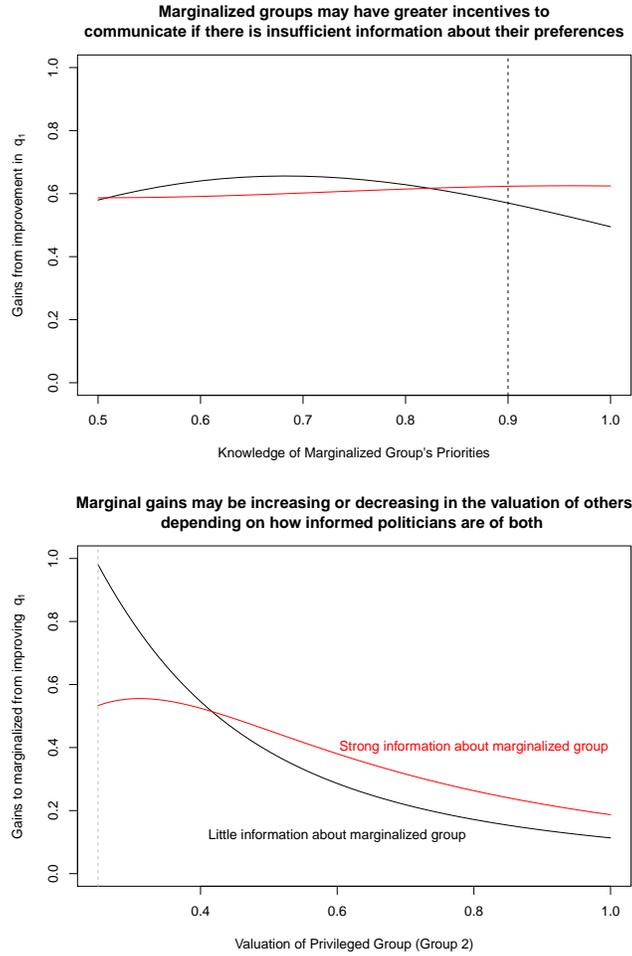


Figure 1: Upper panel: Marginal gains from increasing information about marginalized group’s preferences (red) and privileged groups preferences (black), given prior beliefs about privileged group’s preferences of 0.9. Lower panel: Gains to marginalized group depending on how privileged the privileged group is;  $\beta_1$  set at 0.25;  $q_2$  at .75, for  $q_1 = .5$  (black line) and  $q_1 = .9$  (red line).

The exception is when biases are weak and there is already strong information about the marginalized group relative to the privileged group.

The model’s core results are therefore the following: although marginalized voters might benefit less from informed politicians, they can have stronger incentives to make up the information gap; given equal opportunities, marginalized citizens about whom there is weak information can have incentives to inform politicians more than less marginalized citizens, even though informed politicians prefer to allocate to less marginalized citizens. The incentive to make up the gap comes from two sources. First there are greater effects of information at low levels; second there are strategic incentives: gains from providing information can be greater when politicians are more informed about other groups.

We emphasize that while these logics are consistent with optimal behavior, the results are sensitive to multiple features of this specification, such as the degree of concavity, order of play, and differential costs of access. For example, with log utility, politician allocations do not depend on information. We see this as the key payoff from formalizing our theoretical framework: although the model clarifies the types of logics in operation, it

also makes clear that one cannot expect these logics to hold universally and underscores the importance of empirical analysis.

Several past studies confirm that politicians, at least in many parts of Africa, have limited information on the preferences and priorities of their constituents (Bleck and Walle 2013). Indeed, [authors] report that Members of Parliament in Uganda who were surveyed in-person, describe themselves as insufficiently informed when they vote in plenary and in committee meetings. Weak political communication channels are one reason that, notwithstanding two decades of democratization processes, citizens across Africa generally feel unrepresented, commonly expressing low levels of political efficacy (Lynch and Crawford 2011).

It is interesting to note that, to date, the burgeoning literature on the relationship between information and accountability almost exclusively focuses on the (dearth of) information that citizens have about politicians (Ashworth 2012; Dunning et al. 2018). This study focuses instead on the fact that representatives cannot represent if they lack reliable information on their constituents' preferences and priorities. Indeed, our research project is premised on the idea that improving the information that politicians have—and citizens' awareness that politicians possess such knowledge—may be just as important as improving the information in the hands of voters.

However, studying the logic of political communication described in the model above, and especially the conditions under which different constituents are more likely to invest in communicating their preferences to politicians, is hard. This is because there are likely unobserved characteristics at the individual and the constituency level that are correlated both with the *availability* of communication channels and *intensity* of political communication. We address this challenge by using a field experiment research design in which the availability of a new innovative communication channel has been randomly assigned.

## 2.2 Is Mobile Phone Penetration a Disruptive Technology?

Our study contributes to a growing literature that explores whether the rapid penetration of mobile technology across the developing world can be harnessed to improve governance outcomes (Peixoto and Sifry 2017). Focusing on weak channels of political communication, we ask: can technological innovations that reduce the costs of access to politicians improve channels of political communication, thereby altering MP-constituency relations?

Specifically, we seek to test whether technological platforms that connect voters and MPs, using simple innovations such as text-messaging, can alter the nature of political participation, and ultimately strengthen political representation. Our study is therefore designed to produce evidence that helps us understand how mobile technology might affect who gets to be heard and what gets communicated to political elites. Building on the logic of our model, we assess five dimensions related to this question.

First we look at the overall level of political engagement, by reporting the extent to which citizens adopt a newly introduced SMS-based communication platform to articulate their priorities and preferences (system uptake). Naturally, ICT platforms can be an engine of political change only if there exists an underlying demand to communicate preferences to their representatives.

Second, the extent to which technology can have a transformative effect on politics depends in part on users' identity. Following our theoretical framework, we are especially concerned with uptake among marginalized citizens (i.e., with the flattening of political access). “Flattening” occurs when the share of marginalized

voters using SMS-enabled political communication is high, relative to elite users and relative to the rates at which they use traditional forms of engagement. ICT platforms, on the other hand, would not have a flattening effect if they are used primarily by citizens who are already engaged in frequent communication with political elites using traditional forms of engagement. Flattening is a possible response to opportunities, as shown in the model above, though it is not guaranteed: in some ranges privileged groups about whom politicians are better informed have stronger incentives to render politicians still more informed.

Third, our research is designed to explore the implications of system design choices. Specifically, given our interest in uptake across levels of marginalization, we test whether the cost of communication (price) effects the decision to send text-messages. Price is, indeed, a key concern: the cost of sending a message can affect who communicates and what gets communicated. If messaging is an ordinary good, subsidization will increase the volume of communication that reaches political elites. If marginalized voters are, on average, more sensitive to price, then ICT platforms will increase the relative share of previously excluded groups among system users when offered at lower costs. Alternatively, we could find a larger share of marginalized voters among system users when prices are high, if marginalized voters place a higher value on new access to political elites and are less price sensitive. More advantaged individuals may be more sensitive to prices if they can switch to more traditional channels of political communication. We therefore randomize the cost of messaging and estimate price effects on the overall level of uptake and on the type of users.

Fourth, we explore *feedback* effects—the extent to which knowledge of the use of the system by others affects uptake. If potential users view messages as complements—the more people raise an issue, the greater the likelihood that this issue would be addressed—then feedback effects should be positive.<sup>5</sup> On the other hand, if potential users view messages as substitutes, then feedback would exacerbate a collective action problem, leading to overall lower uptake. In our experiment, we randomize the level of information on system usage and estimate the effect of such feedback on the level of uptake.

Finally, we explore *downstream* effects—the extent to which politicians’ attitudes and behavior might be affected by exposure to citizens articulating their preferences and priorities via text-messaging. Closely related, we also investigate whether citizens’ attitudes—especially their sense of efficacy and trust in government—changes when new political communication channels become readily available.

### 3 Research Design

The field experiment we study was part of the national strategy of Uganda’s Parliament for widening citizen voice. To the best of our knowledge, it is one of the largest political field experiments ever to be undertaken with consenting political elites.<sup>6</sup> Below we describe the political context that gave rise to this intervention—summarizing results from the framed field experiment implemented prior to the national intervention—and describe the design of the national intervention and the data used to study it.

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<sup>5</sup>See Ferrali et al. (2018) who explicitly model messaging politicians as subjected to positive externalities, which is appropriate when feedback can facilitate voter coordination (Arias et al. 2017).

<sup>6</sup>Our study joins a growing body of work using politicians as experimental subjects. See for example, Sheffer et al. (2018) and LeVeck et al. (2014) on politicians’ decision making, and Grossman and Michelitch (forthcoming) for politicians’ response to disseminating information on their performance in office.

### 3.1 Political Context

Uganda provides a good context for exploring changes to behavior in the wake of introducing a new political communication platform. First, Uganda shares characteristics with many low-income countries on relevant dimensions. It is in the mid-range of the World Bank’s low-income economies in terms of economic development (as captured by GDP per capita) and of human development (as captured by HDI ranking).<sup>7</sup> In addition, Uganda is in the middle range ICT ownership, use and access among African countries (World Bank 2016). These factors strengthen confidence in the external validity of our results.

Second, data from Uganda supports the assumption of weak political communication channels leading to dearth of information in the hands of politicians. Consider results culled from a survey the research team conducted with Ugandan Members of Parliament at baseline. We find that the majority of surveyed MPs describe themselves as feeling insufficiently informed when they vote in plenary and in committee meetings. In other work, surveyed Ugandans report that elected politicians do not frequently elicit voter opinions (Grossman, Michelitch, and Santamaria-Monturiol 2017). This evidence suggests that the context is one in which there is an unmet demand for greater information.

Third, results from the framed field experiment (FFE) conducted prior to the launch of the national field experiment, further point to Uganda as a good context for studying the questions at hand. Specifically, findings from the framed field experiment suggests that not only does there exist underlying demand in Uganda for contacting one’s MP via a text-messaging platform, but also that IT communications do not necessarily widen the participation gap between more and less marginalized populations. We briefly describe the FFE below.<sup>8</sup>

The framed field experiment, undertaken in 2011, was delivered alongside a survey conducted in every parliamentary constituency in Uganda using a national representative sample. The FFE sought to assess whether demand existed and to explore the validity of the concern that IT-based communication platforms exacerbate existing inequalities in political access. At the end of the survey, sampled respondents were invited to send a text message to their MP at randomly assigned prices. Discussed in more detail in [citation omitted], the uptake recorded in the FFE—about 5%—suggests that a sizable number of citizens value the opportunity to contact their MPs via SMS.

In addition, usage rates in the FFE were no lower among more marginalized populations, possibly reflecting the fact that these populations have fewer opportunities to access politicians and therefore place a higher value on impersonal and inexpensive ICT channels. Experimentally manipulating the price of sending a text message to one’s MP, we further found, as expected, that reducing the cost of communication encouraged usage.<sup>9</sup> Moreover, consistent with the idea that marginalized populations place a higher value on cheap impersonal communication, we found that marginalized populations were not more sensitive to the cost of political communication than less marginalized populations.

The FFE confirmed that Uganda offers a good context to examine the implications of harnessing technological innovations to improve political communication, and that ICT platforms have a genuine potential to alter citizen-MP relations and “flatten” political access. However, the setup of the FFE also had some limitations.

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<sup>7</sup>Low-human development countries are ranked between 148 (Swaziland) and 188 (Central African Republic). Uganda is ranked 163 (in 2016).

<sup>8</sup>A more detailed description of the FFE can be found in (citation Omitted).

<sup>9</sup>Uptake was almost 50% higher for those randomly assigned to a free SMS treatment arm, as compared to those assigned to a treatment group that was not offered any subsidy for texting their MP.

For example, it allowed only a ‘one-shot’ opportunity to communicate with MPs, and thus was unable to examine usage patterns over time, in which citizens’ behavior is (also) a function of both the usage of *other* citizens and the response of their MP to past messages. Moreover, it was implemented in the context of an in-person survey in which subjects interacted with enumerators regarding their political views. While such personal interaction ensures that subjects are aware of the program, it is also prohibitively costly. Thus, there is no guarantee that using mass communication channels at scale (such as radio) would result in a strong first-stage; i.e., wide-range program awareness. The personal interaction with enumerators may have also made politics more salient to interviewed subjects, further strengthening the invitation to use the platform. The personalized invitation to contact one’s MP may have also increased both the sense of empowerment and civic obligation to raise one’s voice. It is also possible that subjects perceived the FFE as closer to a civil society effort than an official government program.

These considerations raise the question of whether similar effects would be found when the ICT service was brought to scale, and shifted from being a researcher-led initiative to being an institutionalized part of national politics. The field experiment described in the next section was designed to address these concerns.

## 3.2 Intervention

As part of the Ugandan Parliament’s national strategy, a case management platform hosted in the National Parliament was developed, allowing citizens to send messages to their MP via SMS or a voice call to a call center. MPs randomly assigned to participate in the program (“uSpeak”) were given access to the platform and trained in its use. The platform allowed MPs to log onto a dashboard where they could read tagged SMS messages from constituents, reply, and see simple descriptive statistics about the messages they received, such as what the priority issues in their constituency were within a selected time-frame. A screenshot of the query dashboard is presented in the Supplementary Material (Figure 1). Only treated MPs were able to receive messages from their constituents via the case management system.

The ICT platform was promoted to citizens through 30 seconds radio advertisement spots, played twice daily on local radio stations over the study’s six-months period. The radio ads were in local languages, and featured a skit where actors portraying constituents talked about how uSpeak could be used to draw the MP’s attention to important issues, specifically service delivery deficiencies. These skits were first tested using focus groups. A second tier of randomly assigned treatments—price and feedback—was also delivered via the radio ads.

### Treatment 1: Elite Participation

The NFE involved 186 MPs who volunteered to be part of a six-month pilot. It was expected that, if deemed successful, all MPs would be phased into the program at the end of the study. Given the sensitivities of providing a new service to only some constituencies, it was agreed that MPs would be selected into the program using a public lottery managed by NDI. Block randomization was used to assign MPs to treatment groups; MPs were sorted into bins based on their *type* (Woman MP or Constituency MP), *party*, and *region*.<sup>10</sup>

### Treatment 2: Variation in Price

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<sup>10</sup>Each bin was used to implement a separate public lottery with a target number of MPs selected into treatment based on that MP type’s prevalence in the subject pool. Block randomization was used not simply to improve balance in expectation, but also to improve ex-post equality between parties in participation.

To assess the effects of price on uptake, Parliament randomly varied the cost of sending a message to MPs via the uSpeak system, across and within constituencies. Each constituency was assigned 3 months in which uSpeak would be provided free of charge and 3 months without any subsidization. Being sensitive to potential sequence effects, all possible sequences of full price and free months were randomly assigned to constituencies in the treatment group using a blocked design. Note that while the variation in prices in the first period provides a clean separation into price groups, for identification based on variation in subsequent months we must assume no carryover effects.

### **Treatment 3: Variation in Feedback**

In order to examine whether information on *others'* usage encourages greater uptake, we added a 'feedback' treatment arm delivered through modification of the base radio ads. In one version, voters heard that others had been sending messages to the system about the need to do more in the educational sector. A second variation also highlighted the educational sector but without communicating that others had been using the system to lobby in that area. To the extent that there are complementarities in public goods messaging, we expect that hearing that others are sending messages about education should increase the willingness to contact one's MP. Indeed, our feedback skit was written explicitly in a way that made this sort of complementarity more apparent to radio listeners.<sup>11</sup>

We selected eight unique price sequences and six unique combinations of the feedback treatment that together produced 48 unique combinations of price and feedback sequences. These were assigned in a balanced way to treatment constituencies, resulting in roughly two constituencies of each unique treatment schedule. In Supplementary Materials (Figure 6) we provide an example of treatment schemes for a subset of constituencies.

## **3.3 Data**

Data for testing the effects of the uSpeak program come from four sources: (1) a baseline survey of Ugandan adults randomly drawn from all constituencies in Uganda, conducted immediately following the 2011 Parliamentary election, (2) the SMS messages sent by constituents to the uSpeak system, tagged with the date and time they were received, (3) a callback phone survey we conducted with uSpeak users, and (4) an endline survey of a nationally-representative sample of Ugandan adults in a subset of the uSpeak constituencies. In addition, as described below, we conducted a follow-up experiment with about 3,000 Arua district residents to help adjudicate some of the conflicting findings between the natural field experiment and the framed field experiment.

## **4 Main Results**

We focus on core results related to overall uptake, flattening (the characteristics of participating populations), price and feedback effects, and downstream effects. We note that uptake and flattening are not experimental treatment effects in the usual sense, rather they are levels assessed under controlled conditions. Price and feedback effects draw on randomized variation within treatment and downstream effects draw on randomized

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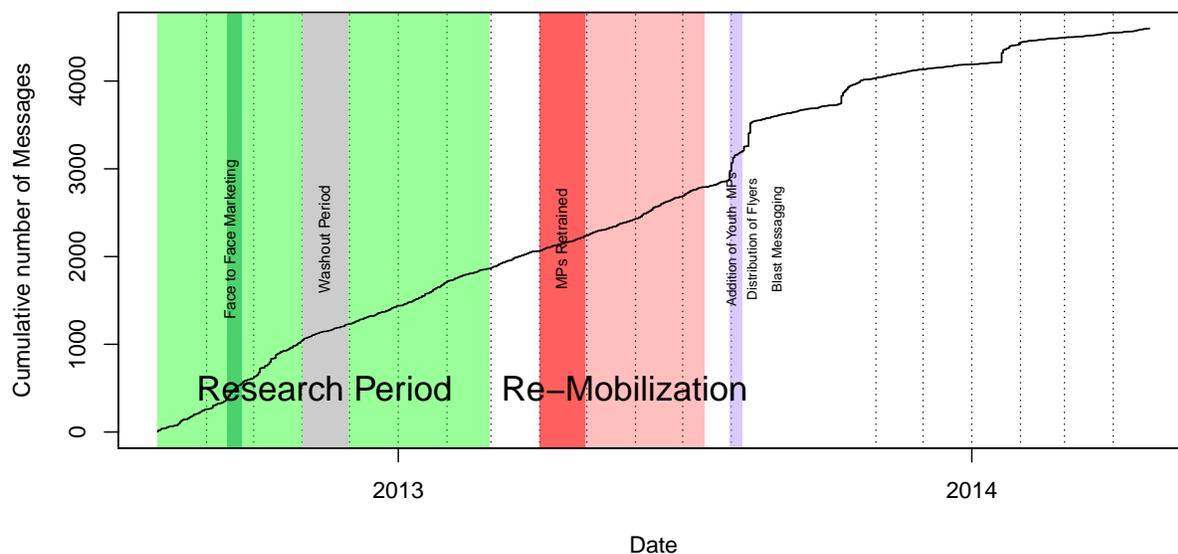
<sup>11</sup>By contrast, if people view text messages as substitutes, then hearing that others are using the IT system could exacerbate the collective action problem.

MP participation in the intervention, as described above. Analyses implemented to explain our results on uptake and flattening are described in Section 5.

## 4.1 Weak Uptake

Unlike the FFE described above, uptake in the NFE was very low. Despite twice daily radio ads and price subsidization throughout the country, MPs in the treatment group received a total of 1946 messages during the 6-month study period. Using the most recent 2014 population census, we estimate conservatively that the radio ads were played over an area where 10 million voters live. This uptake then corresponds to a monthly usage of about 1 in 30,000. Figure 2 shows the cumulative messaging over time, extending beyond the study period to show uptake in the post-study period including various periods in which an assortment of mobilization efforts were used by Parliament and NDI—none of which produced sustained effects (see also Supplementary Materials, Figure 7).

Figure 2: USpeak Natural Field Experiment: Uptake



**Note:** Cumulative messaging over time. Gray area represents the wash-out period in which no radio spots were played. Green areas denote the period with experimental variation. The figure also shows uptake in the post-experimental period, in which there were attempts by Parliament and NDI to further encourage usage.

A broad categorization of the types of messages suggests that, as with the Framed Field Experiment (reported in [authors]), a large share of messages were for local public goods or local community interests with a much smaller set for national or policy concerns; a much larger share of messages here were of a more personal nature, accounting for nearly a half of messages sent compared to at most 10% in the FFE. See Supplementary Materials, Table 4.

## 4.2 No flattening effects

One of the key findings of the FFE was that the share of marginalized populations—such as women and the poor—among system users was higher than the share of marginalized constituents participating in traditional forms of political engagement. That finding formed the basis of our conclusion that ICT platforms have a genuine potential to flatten political access.

To assess flattening in the national experiment, we conducted a phone survey of system users. Using a call center that the research team had set up, local enumerators contacted all uSpeak users no longer than two months after they had sent a text message to their MP. The short callback survey was designed to elicit information on users’ demographics, on whether they received a response from their MP, and general satisfaction with the ICT service.

Comparing results from our callback survey to information culled from the FFE, it is clear that the scaled-up national program failed to replicate the flattening effect identified in the FFE. Specifically, the users of the uSpeak system were wealthier, more highly educated, and overwhelmingly male, compared to those sending text-messages in the FFE. Put plainly, the uSpeak program failed to elicit participation from marginalized populations in the way political actors expected. Figure 3 provides information on the distribution of wealth, gender, and education, across the two field experiments.

## 4.3 Insensitivity to Price

Unlike the FFE, we find no evidence of overall sensitivity to price in the scaled-up national program. Monthly rate of messaging in the free and full-price treatment conditions are almost indistinguishable (see Figure 7 in Supplementary Material). Testing for a price effect more formally, we run a linear regression of the number of messages received in a given month on price—a binary variable that takes the value of one for full price and zero for months of free messaging—controlling for the month feedback treatment indicator and MPs fixed effects. Results presented in Table 1 suggests that contrary to the FFE, in the scaled national program, price did not significantly affect uptake.

The lack of evidence of a price effect could be explained by the fact that, as we mentioned above, those who chose to send an SMS to their MPs via the ICT system were a small number of politically engaged citizens from, by and large, traditionally powerful groups. Given that at the time of the study the average cost of text-messages was lower than 110 Ugandan Shillings (equivalent to about 2 US cents), it is reasonable to interpret the null effect of the price treatment as stemming from the fact that relatively well-off citizens are not price sensitive when communicating directly with MPs.<sup>12</sup>

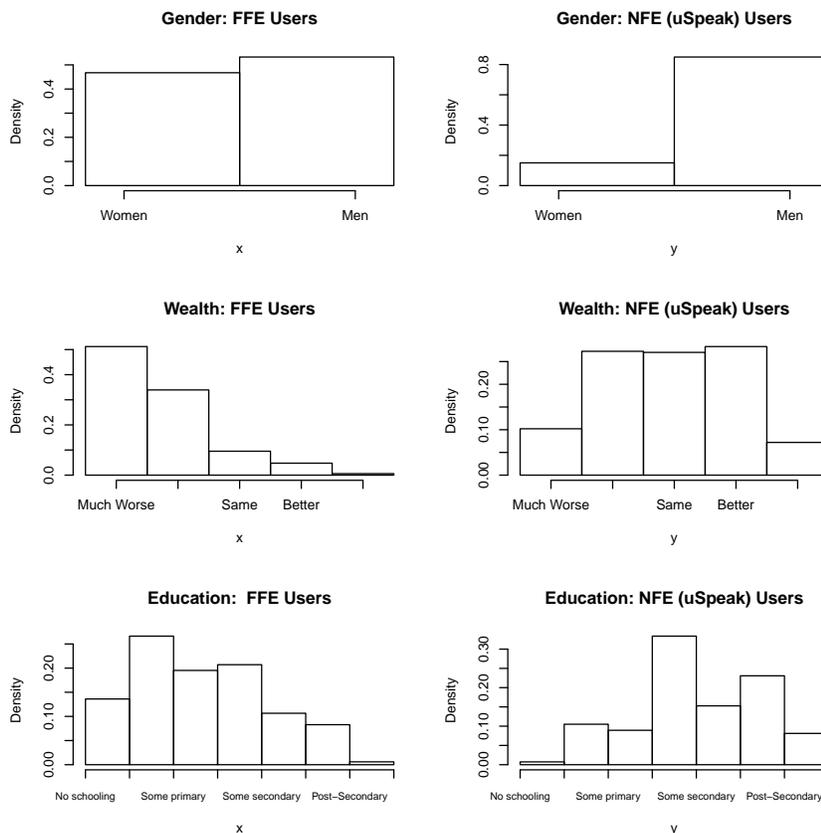
## 4.4 No Evidence of Downstream Effects

Thus far we have shown that uptake in the scaled-up uSpeak program was low and that fully subsidizing the cost of messaging did not increase voters’ proclivity to contact their MP via SMS. Notwithstanding the low rates of usage, it is possible that uSpeak had a positive effect on voters’ sense of efficacy and their satisfaction with politics in Uganda. This would be the case if citizens view the existence of the ICT platforms,

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<sup>12</sup>The divergence observed in price effect across experiments is reminiscent of the way subjects of controlled laboratory experiments react to even small monetary manipulations that are inconsequential outside laboratory settings.

Figure 3: Demographic Differences: Users in the Framed Field Experiment compared to users in the Natural Field Experiment



**Note:** Users in the scaled-up NFE were more likely to be male, better educated, and wealthier than users in the FFE. Data Sources: phone surveys of all system users.

irrespective of one’s own usage, as an important tool for strengthening citizen voice. This was a goal of the intervention and we report on it here briefly. Results in this section use experimental estimates of the effects of the intervention, exploiting the random assignment of the scaled-up program.

To test for the effect of the national program on voters’ efficacy we turn to our endline survey. The survey, which took place in July-August 2014, included 2,714 adult respondents from 76 constituencies and 304 villages in 52 districts across Uganda.

To measure efficacy, we asked survey respondents whether they agree with the following statement: *People like you can do things that can have an influence on the actions of . . . [your constituency MP]*; we then repeated the question for the president, district chair, and traditional leaders, which serve as placebo tests. Our key dependent variable is a binary indicator that is equal to one for the 60% of respondents who had agreed that citizen action could influence their MP. We then run a simple OLS model regressing the efficacy outcome on a treatment indicator and district fixed effects. Results, presented in Figure 4, suggest that uSpeak had no discernible effect on voters’ sense of efficacy. Note that the graph also gives results from four placebo tests, assessing increased confidence in leaders that are not related to uSpeak, and, surprisingly, passes two of these. While surprising, the pattern suggests that the intervention did not increase the efficacy

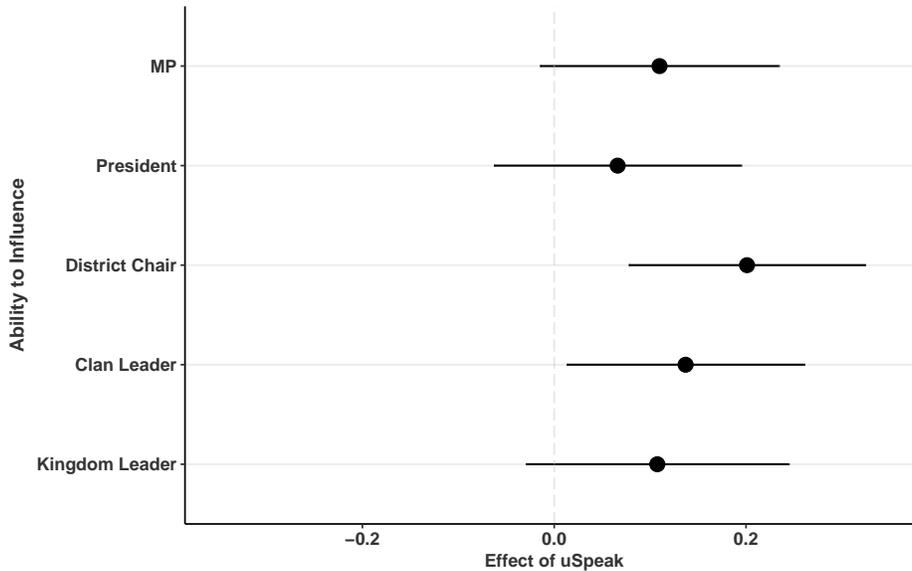
Table 1: Uptake as a function of price and feedback

	<i>Dependent variable:</i>	
	Messages per month	
	(1)	(2)
Price	-0.081 (0.262)	-0.079 (0.262)
Education prompt		-0.275 (0.453)
Education plus Feedback Prompt		-0.132 (0.452)
Observations	660	660
R <sup>2</sup>	0.055	0.055
Adjusted R <sup>2</sup>	-0.145	-0.149
F Statistic	5.251*** (df = 6; 544)	3.972*** (df = 8; 542)

*Note:* \*p<0.1; \*\*p<0.05; \*\*\*p<0.01

of citizens with respect to MP engagement relative to the effect on engagement with other political actors.

Figure 4: Efficacy Effects



**Note:** The marginal effect of uSpeak on political efficacy measured as respondents' perception of their ability to impact their MP.

## 5 Discussion

Experimental findings from the national program conflict with the results from the FFE. Notably, uSpeak resulted in low uptake, even when the service was offered to voters at no cost. Moreover, confirming concerns that ICTs would exacerbate existing inequalities in political access, when uSpeak was used, it was by and large used by citizens whose voice is already more likely to be heard. In other words, the groups that have the weakest access to political processes were also the least likely use the new ICT platform. We first explore some of the reasons that may account for uSpeak’s low uptake, we then assess several explanations for the fact that—contrary to the FFE—marginalized populations were significantly less likely to use the new ICT platform.<sup>13</sup> Our goal here is not merely to account for these diverging results, but rather to use the analysis to derive substantive insights regarding the role ICTs can currently play in improving political communication in low-income countries.

### 5.1 Explaining Low Uptake

Although the FFE led by the research team was meant to capture the key features of the scaled-up national ICT platform, the introduction of relatively tight experimental control introduces a number of differences.

We first explore the explanatory power of two external features of the NFE—which are common to interventions that are scaled-up from controlled pilots to larger-scale programs—that may have been consequential. We refer to these as “scale” and “agent” effects. In addition, we examine the implications of subtle differences in the delivery of the treatment. These “design” effects may be especially relevant for interventions that involve the dissemination of information to subjects.

Changes in *scale* are often described as a problem of general equilibrium effects (Deaton 2010). This concern is of particular salience when treatment effects are sensitive to the share of treated in the population. Scale effects are of special concern when subjects can accurately infer the magnitude of a program from its delivery method, as is clearly the case in our study. In our setting, it is quite possible that collective action problems get altered substantially as scale increases. Insofar as political communications complement each other, or substitute for each other, increases in scale could lead to greater or lower overall levels of communication.

A second possible reason for the low uptake relates to *agents*. Whereas the research team implemented the FFE, the Parliament of Uganda and NDI led the scaled-up national program.<sup>14</sup> In our case, this change in agents might have affected citizen expectations regarding the responsiveness to their messages. In other words, the fact that the scaled-up national intervention was implemented by Parliament rather than by researchers may have reduced the incentives of the target populations to engage.

The third possibility relates to *experimental design* and specifically, to the possibility that details of the mode of treatment delivery—the nuts and bolts of executing field experiments—mattered a great deal for citizens in deciding whether or not to communicate with elites. We focus here on two possibilities. The first is that the method of delivery (radio spots) introduced a *treatment compliance effect*: that Ugandans were simply unlikely to hear or internalize appeals issued through mass media, and not less likely to respond, conditional

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<sup>13</sup>We do not explore the reasons behind the lack of downstream effects, since the fact that the scaled-up national program generated such weak first-stage results speaks volume to why voters and politicians’ attitudes and behavior were not affected by the introduction of uSpeak.

<sup>14</sup>Differences in agents across scales are common: for example, the Millennium Villages initiative sought to assess the scope for government led development change by examining an intervention in which government was not a primary actor.

on hearing. This mode of delivery differs from the FFE where the enumerators ensured that respondents unambiguously heard and internalized the information on the SMS platform.

Closely related is the possibility that different methods of delivery have varying degree of an (implicit) *invitational effect*. It is possible that uptake was relatively high in the FFE not simply because the in-person survey context ensured awareness of the new ICT platform, but also because the enumerators had personally invited respondents to contact their MP. Communicated in the context of a survey, such invitation may appear as a more personal encouragement to engage in politics. A direct personal invitation has an empowering effect, signaling receptiveness and the possibility that political communication will make a difference. These last two ‘design’ mechanisms are closely related yet distinct: one is about whether an invitation was empowering and deemed personal, the other is whether an invitation gets heard at all.

We use a number of strategies to adjudicate between these four explanations. First, we exploit a feature of the scaled-up field experiment in which there was variation in the *feedback* provided to voters on the behavior of others. This allows us to examine whether exacerbating collective action problems due to scale can, at least partially, account for the significantly lower uptake. Second, we conducted a citizen endline survey with a nationally representative sample, which allows us to assess—albeit with some lag—ex-post differences in treatment compliance. Third, we hired a Ugandan private marketing firm to examine whether the radio stations NDI had contracted indeed played the ads according to the experimental design. Fourth, we implemented an additional “mechanism experiment” in one district in Uganda in which we specifically varied the invitational component. Table 2 summarizes the list of potential explanations, and the source that was used to explore their explanatory power.

Table 2: Differences between the FFE and the NFE

	Type	Data Source
1	Scale Effects	National experiment (Feedback treatment)
2	Agents Effects	Users survey data (callback)
3a	Design Effects I: Treatment Compliance	Citizens survey data (national sample), Radio monitoring data
3b	Design Effects II: Invitational Effects	Follow-up Mechanism experiment

### Low Uptake: Scale effects

Political communication is subject to collective action problems. If many others are visibly lobbying a politician, free-ridership may become more likely if messages work as substitutes. By contrast, visibility of lobbying can improve voters’ coordination leading to a cascade of usage, assuming a sufficient number of voters view messages as complements.

From the model in Section 2.1 above we can see two ways in which a logic of this form can play out. First, if politicians are more informed about others, and better able to target resources to them, this can increase incentives to provide a politician with information on one’s own preferences. Second, scale may also result in lower individual contributions through a simple logic of substitution for members of a given group. In the extreme case in which information from constituency members were perfect substitutes and citizens faced linear costs, increases in potential information providers would not alter the amount of information provided, in equilibrium, which in turn implies a corresponding reduction in per capital information provision.<sup>15</sup>

<sup>15</sup>More formally, say, utility for an agent  $h$  in group  $i$  were given by  $u_h(\alpha_i(q_i(\sum_{k \in i} a_k))) - a_h$ , keeping the politician’s

Conversely, one could also construct examples in which when many others are lobbying for a common good there may be increasing returns to lobbying (or here too, there may also be increasing incentives to free ride). In short, if the incentives to use technological innovations for political communication depend on the perception of how others are engaging (Ferrali et al. 2018), then outcomes at a small scale may look very different to outcomes at a large scale. From this perspective, weaker participation from the scaled-up program may reflect a simple failure of collective action.

In order to understand if changes in scale induced free-riding, we look for differences in outcomes due to our *feedback* treatment. Recall that in the scaled up national experiment, we exogenously varied the information constituents received about the level of activity by other voters in previous periods. In particular, a random subset of constituencies was informed, through the short radio ads, that other voters had been using the system to mainly raise issues around education. Under a free-riding logic, such information would depress engagement among those exposed to it.

Returning to Table 1, however, we find no evidence of sensitivity of engagement to information on uptake by others—neither the difference between feedback on education messages and standard marketing spots, nor differences between education marketing with and without feedback is significant. This is consistent with a set of analyses we conducted on the data from the FFE in which we found no evidence for strategic engagement with the system. We conclude that scale, by itself, does not seem to be a key factor driving our divergent results in terms of overall uptake.

### **Low Uptake: Agent Effects**

Another possible reason for the low uptake witnessed in the scaled-up national program is *agent effects*. Citizens’ usage of mobile messaging plausibly increases with the belief that there is a receptive representative at the other side of the interaction (Grossman, Michelitch, and Santamaria-Monturiol 2017). In the model in Section 2.1, these effects are captured by the  $\beta$  terms; citizens will be less likely to communicate if they think that a political actor cares less about their interests. More subtly, they will also be less likely to participate if political actors are better informed about the interests of rival constituents.

Which system should voters expect to produce greater responsiveness by politicians? On one hand, unlike our FFE, the scaled up national program is formally owned and led by Parliament, which signals some level of commitment by politicians. In addition, the dynamic nature of the scaled-up program—i.e., the ability of MPs to interact with citizens directly via the ICT platform—further allows MPs to signal their responsiveness directly. This sort of dynamic reciprocal relationship could not have been established in the ‘one-shot’ controlled FFE. On the other hand, in the scaled-up program, the communication between citizens and politicians was direct, whereas in the FFE this relationship was mediated by the research team, which was responsible for delivering the messages to survey respondents’ respective MPs. Citizens may believe that their MP will take their messages more seriously if researchers or an NGO mediates the relationship between voters and representatives; for example, if it follows up in case some messages get ignored. Thus it is hard to predict *a priori* how the change in the implementer’s identity would affect citizen uptake. We explore (non-experimentally) agent effects in two ways.

First, we use the callback survey ( $n = 2517$  uSpeak users) to calculate MPs’ response rate at the constituency

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information about other groups and bias fixed. Here  $q_i$  (we ignore the sector superscript) is replaced by the function  $q(\sum_{k \in i} a_k)$  where  $a_k$  is an information-providing action of an individual in group  $i$ , taken here to be substitutes. In this case, in any interior solution, the individual information contributions sum to a constant, no matter how many potential information providers there are in a group.

level and then test for a correlation between MP’s responsiveness and the volume of messaging at the constituency level. We find that only 9 percent of uSpeak users report ever hearing back from their MP; in fact, in almost half of the treated constituencies (44) *not a single uSpeak user had received any response from their MP*. Moreover, analyzing system login information, we find that the majority of MPs did not even read many of the messages sent to them. As expected, we find a positive correlation between messaging and responsiveness, which is consistent with citizens’ low engagement being a rational response to their MP’s (in) action during the scaled-up study period.

Second, though the callback survey analysis focuses on system users—a self-selected group—we, nonetheless, can also assess whether broader expectations regarding MP inaction may have contributed to the low uptake rate among the general population. Here we examine responses in the citizen endline survey, when our national representative sample was asked to indicate reasons for why people might not use SMS platforms such as uSpeak to communicate with their MPs. Figure 5 (top left) provides information on the share of respondents in treatment constituencies that indicate each possible reason. Tellingly, we find that close to 50 percent of respondents, report that they would not send a message because they do not expect their MP to be responsive, and about a quarter report a reluctance to contact their MP via text-messaging out of fear of bad repercussions.

We do not have information on the expectations of responsiveness from MPs in the FFE and so cannot compare those directly. Nevertheless, the statements by citizens and the very weak responsiveness by politicians suggests that the low engagement with the scaled up program was a rational response on the part of citizens.

### **Low uptake: Treatment compliance**

It is possible that the difference between the NFE and the FFE is simply due to a weak first-stage; i.e., that the radio ads had a limited reach and thus an overwhelming majority of constituents never heard of the uSpeak program. To test for *compliance* effects we asked respondents in our citizen endline survey directly whether they have ever heard about uSpeak. The survey was implemented in 50 parliamentary constituencies approximately one year after the six-months radio campaign, though at a time when the uSpeak system was still active and promoted by Parliament. In light of the time gap, we used a deliberately strong prime, which entailed playing the original radio ad and asking respondents if they have heard of the service (uSpeak) the ad sought to promote.

Starting with the raw data, we find that about 17% of respondents in control constituencies and 24% of respondents in treatment areas self-report that they ever heard of uSpeak.<sup>16</sup> Note that control respondents are not necessarily misrepresenting their knowledge of the program; this is because radio signals normally have a range that encompass more than a single parliamentary constituency.<sup>17</sup> Testing for a first-stage more formally, we take a conservative approach: first, we code respondents that live in constituency  $j$  as treated if either their constituency (usually male) MP or their district (women) MP were assigned to the uSpeak program, and then calculate the share of the constituency hearing about uSpeak adjusting using survey sample weights. Regressions results at the constituency level, using inverse propensity weights based on both constituency and district assignment probabilities, are reported in Table 3. We find a large, positive and significant first-stage. This result is robust to whether or not we control for (aggregated) individual-level

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<sup>16</sup>We note that the reported rate of “compliance” likely is an upper bound due to the possibility of social desirability bias.

<sup>17</sup>Hearing the message in control areas does not imply non-compliance since the ads were tailored to employ the name of treated MPs only. Control subjects could be aware that others were treated but this does not make it possible for them to take-up the treatment.

covariates and whether or not we add fixed effects for the randomization stratification blocks.

Table 3: First Stage for Constituency MPs only

<i>Dependent variable:</i>				
	Knowledge of uSpeak			
	(1)	(2)	(3)	(4)
Treatment	0.125*** (0.022)	0.12*** (0.022)	0.105*** (0.032)	0.095*** (0.034)
Education		0.026 (0.038)		0.058 (0.043)
Income		-0.011 (0.033)		-0.036 (0.035)
Constant	0.089*** (0.013)	0.08 (0.067)		
Block FEs	no	no	yes	yes
N	50	50	50	50
<i>Note:</i>	*p<0.1; **p<0.05; ***p<0.01			

Second, we estimate constituency-level SMS sending rates (take-up) as a function of hearing about uSpeak using two stage least square regressions. This analysis allows us to compare directly the take-up rate conditional on hearing about the program in the NFE against those of the FFE; it also serves as a reality check for the first-stage regression since the dependent variable here is a behavioral measure derived from the uSpeak database. As reported in Table 4, hearing about uSpeak has a take-up rate of 0.002 with standard error of 0.001; in other words, the take-up rate is between 1/4 and 1/10 of 1 percent. Comparing this to a take-up rate of about 5% in the FFE (which is larger by a factor of about 25), we can confidently reject the null of no difference between NFE take-up among compliers and our estimated FFE take-up.

This analysis has two implications. On the one hand, the effectiveness of radio as a marketing device is not strong. Indeed, when probing deeper about respondents' knowledge of the uSpeak program we find that only 6 percent of treatment respondents were able to confidently say that their MP had participated in the program. Moreover, when asked to repeat the four-numbers short-code, less than half a percent of treated constituencies claimed to know the short-code to send a text-message to their MP and an additional 3% report they once knew the number but have since forgotten it. These findings strongly point to the limitation of radio marketing to garner sufficient awareness to the new service.

On the other hand, this is clearly not the full story. Take-up differential cannot be fully explained simply as a function of a weaker first-stage in the NFE since the two stage analysis suggests that the effects on those that do get the message are much weaker than in the FFE.<sup>18</sup>

### Low Uptake: Invitational Effects

<sup>18</sup>A back of the envelope calculation suggests that if messaging got through to 10% of up to 10 million subjects and had 5% of these responded, there would have been 50,000 messages entering the system.

Table 4: Second Stage for Constituency MPs only

	<i>Dependent variable:</i>	
	(1)	(2)
	Uptake	
Knowledge of uSpeak	0.002*** (0.001)	0.002* (0.001)
Constant	-0.00016** (0)	-0.00027 (0)
Block FEs	no	yes
N	50	50
<i>Note:</i>	*p<0.1; **p<0.05; ***p<0.01	

We turn to explore the possibility of a second design effect; namely, that the marketing tools used to inform citizens about a new service or program likely have (unintentional) invitational effects. Recall that the two experiments differed in their mode of marketing: whereas the scaled-up national program used 30 seconds radio ads, in the FFE, respondents were invited by enumerators to contact their MP in the context of an in-person survey. As mentioned, direct personal invitation may have an empowering effect, or it may signal greater government responsiveness. Multiple logics consistent with the model in Section 2.1 could underpin these effects—an invitation could in principle change a voter’s beliefs about  $\beta_i$ —how much the politician cares about their welfare—as well as about the collection of  $q_{ij}$  parameters—politician’s knowledge of citizens’ preferences. As further discussed below, if such invitational effects operate differently for marginalized and non-marginalized populations, this could account for the differences in observed flattening effects.

To assess the role personal invitations plays in the decision to politically engage using an ICT platform, we implemented a third (mechanism) experiment. To do so, we made use of an existing SMS platform, UBridge, which has been operating in Arua district since late 2014. UBridge was developed in partnership between UNICEF’s Ureport platform and Uganda’s Governance, Accountability, Participation and Performance [GAPP] project.<sup>19</sup> Unlike uSpeak, which connects citizens with national MPs, UBridge was designed to open a new channel of communication from citizens to local government officials to specifically report problems of public service delivery. UBridge was launched as a pilot study in over 100 villages across Arua. A study evaluating the effect of getting access to the UBridge system is underway and is not the subject of this paper. At the time of our ‘mechanism’ experiment, UBridge had 4,568 registered users, out of which 2,720 were explicitly verified by the research team.<sup>20</sup>

On June 13, 2015, UBridge conducted a baseline poll using a robocall system asking users about their attitudes toward budgetary processes. The key outcome of interest is a binary variable that receives the value of 1 if the UBridge user responded to the poll, and 0 otherwise. Of the 2,720 verified users, 12% responded to the opinion poll and shared their views with UBridge. To explore the role of direct invitations on levels of ICT-based political engagement, we asked UBridge to run a modified version of their baseline poll but

<sup>19</sup>With some loss in external validity, our design aims to keep the *treatment compliance* effect constant by focusing on respondents in the UBridge system. We hope that parsing the outcome compliance effect will be the focus of future studies.

<sup>20</sup>We verified the identity of registered users through a call center that we set up with the help of Innovation for Poverty Action, Uganda.

now experimentally introducing a modest variation in their outreach activities. All users would be invited to participate in an opinion poll regarding taxation, similar to the previous UBridge poll. In a randomly selected treatment group, however, UBridge preceded the call with a set of (blast) text-messages that explicitly invited participants to take part in the weekend poll and that highlighted the importance of individual responses. Further details on the block randomization used in this experiment, as well as the full text of the treatment text-messages, are provided in the Supplementary Materials, Section 2.

Our primary measure is the response (or non-response) by UBridge users to the weekend opinion poll. The encouragement text-messages were delivered on 24, 25, and 26 June 2015, and the poll took place on 26 June. We estimate average treatment effects using a regression that accounts for block fixed effects. Our analysis takes account of the variables used for blocking but introduces no further controls. Our primary regression uses only the verified subset of UBridge users, whereas our secondary analysis includes all registered users whether or not they have been positively verified.

Results, reported in Table 5 (column 1), suggest that invitation had a large positive effect on response rate: 2 percentage points from a base rate of 9.4% for the control group. These results are consistent with findings reported by Grossman, Michelitch, and Santamaria-Monturiol (2017) in a similar context, and by Dale and Strauss (2009) and Malhotra et al. (2011) in the USA. We note that even though the invitation tested in the mechanism experiment was relatively weak (three text-messages) compared to the in-person invitation used in the FFE, it was able to increase participation rates by over 20%. The evidence at hand allows us to conclude that, consistent with insights from the voter mobilization literature (e.g., Green and Gerber 2008), more personal invitations can have a powerful effect on rates of participation. We can further conclude that, at least in low-income countries with characteristics similar to Uganda, short radio ads likely represent a marketing strategy that is too impersonal to mobilize large-scale participation.

## 5.2 Flattening: Differential uptake by marginalized groups

We have focused on explaining the overall low uptake witnessed in the scaled-up national experiment, especially when compared to the (relatively) high demand observed in the FFE. We now build on the theoretical framework developed above (and summarized in Table 2) to explore some of the reasons that may account for the lack of “flattening” effect; namely, that when brought to scale, marginalized populations used the ICT platform to articulate interests at significantly lower rates as compared to more powerful populations.

### Scale effects and differential uptake

Returning to the logic described in Section 2.1, scale can affect not simply per capita information provision but also the profiles of providers. Most simply, reduced flattening could result from increases in scale if substitution effects were stronger within more marginalized groups.<sup>21</sup> While a theoretical possibility, the generally weak evidence for any scale effects suggests it is unlikely that scale could explain differences in flattening.

### Differential Uptake: Agent effects

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<sup>21</sup>In the model, marginal gains from providing more information can be weaker for more marginalized groups and in some cases can be weaker the more informed politicians are about other groups. In principle, information provision from other groups may differentially reduce the incentives for marginalized groups to provide information, leading to a reduction in flattening. There are conditions however in which this cross partial can take the opposite sign, resulting in an ambiguous prediction.

It is also theoretically possible that agent effects explain the differences in flattening effects. This could arise if marginalized voters are more doubtful of government willingness to engage and are encouraged more by the presence of NGOs. As NGOs often provide services to marginalized communities where governments fail to provide, this explanation has face validity. The data reported in Figure 5 (top row), do not bear this out however. Although less educated voters were marginally more likely to doubt MP responsiveness, the difference is not large and trends in the opposite direction for poorer citizens.

### **Differential Uptake: Treatment compliance**

Recall that to test for compliance effects, we asked respondents in our endline survey whether they have ever heard about uSpeak. Treatment compliance can help account for differential uptake if more marginalized populations have lower awareness of the new political communication service. Figure 5 thus breaks down the overall awareness of uSpeak by sex, education, and wealth. The figure demonstrates clearly that, marginalized groups, such as women, less-educated, and less well off groups (denoted in black), are significantly less likely to be aware of the uSpeak program, compared to less marginalized groups (denoted in gray). We note that these differences are least pronounced for education.

To further assess whether a treatment compliance effect can be responsible for the large gaps in participation, we asked survey respondents about radio ownership. Starting with gender gaps, we find that 0.71 of female respondents reside in a household that owns a radio in good working condition, compared to 0.76 of men. Since owning and listening to radio are distinct, we further asked our representative sample about their frequency of radio consumption on a five-point scale (from “never” to “daily”).<sup>22</sup> Figure 5, which provides information on the distribution of responses broken down by gender, indicates that women listen to radio much less frequently than men. Moving to another form of marginalization, poverty, we find that poorer respondents were only marginally less likely to have heard of uSpeak though they were significantly less likely to listen to the radio (see also Supplementary Materials, Section 4.2).

While we have demonstrated above that treatment compliance cannot fully account for the difference in take-up rates between the NFE and FFE, the findings reported in Figure 5 provide evidence of treatment compliance effect as one explanation for the relatively low participation rate of marginalized groups in the uSpeak program.

### **Design Effects II: Invitational effects and differential uptake**

In our final analysis we assess whether ‘invitational’ effects can help account for the differential uptake across more or less marginalized groups. The intuition is that personal invitation to participate in politics may have a larger marginal effect for marginalized populations who may have, on average, lower political (internal) efficacy to begin with.

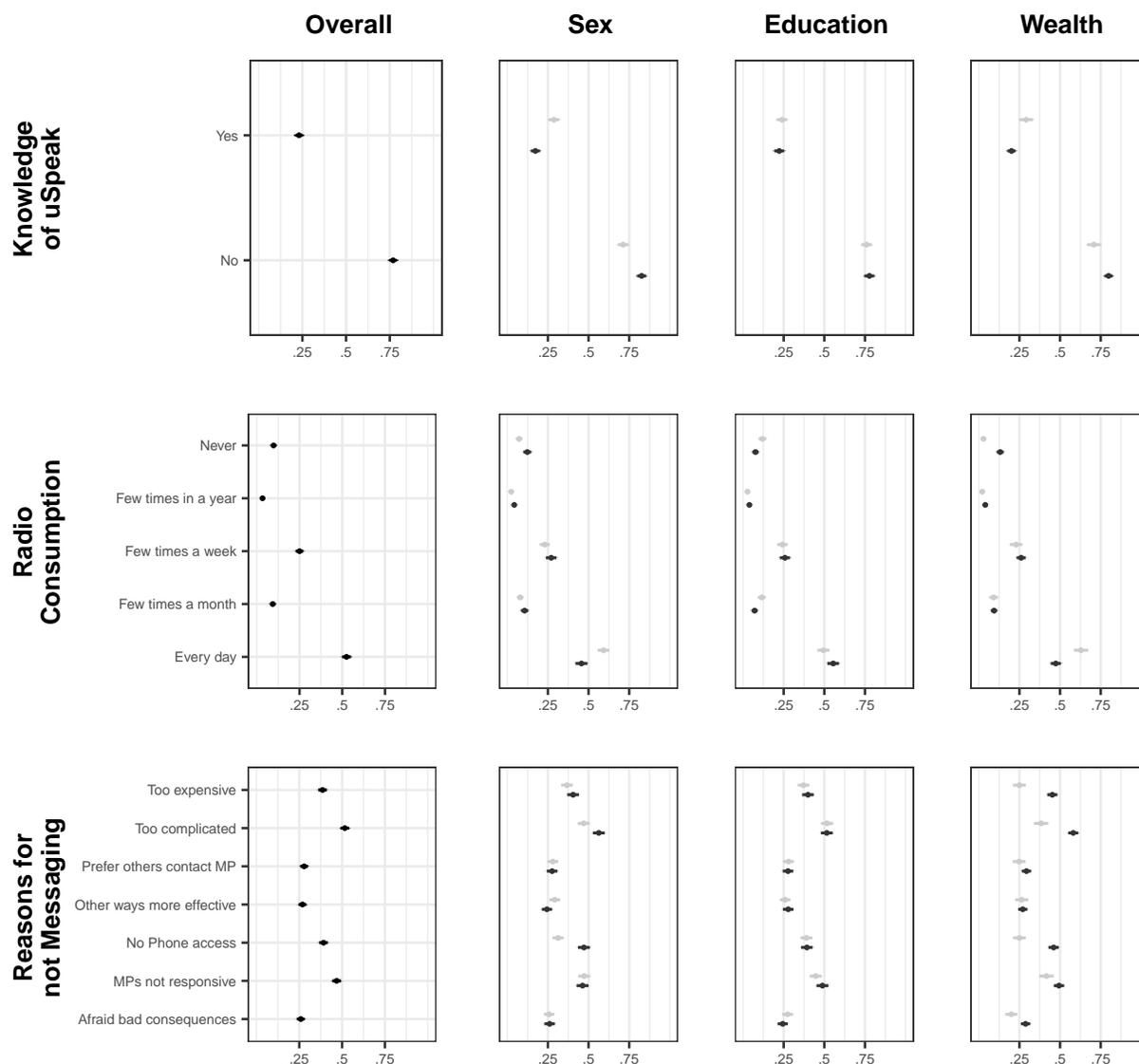
Recall that to assess the role of invitations in political communication we implemented a third “mechanism” experiment, using an existing political IT communication platform, UBridge. Recall also that in this experiment our key outcome of interest is whether UBridge users responded to an opinion poll. It is important to note that in implementing the “mechanism” experiment, we were limited by the amount of information we had on registered UBridge users. Specifically, the only marginalization information the research team had on UBridge users is their gender.

Consistent with past findings demonstrating gender gaps in participation in Africa (Isaksson, Kotsadam,

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<sup>22</sup>The question verbatim was: *How often do you get news from the following sources?*

Figure 5: Possible causes for non-engagement across population categories



Note: Differences between more privileged (grey dots) and more marginalized (black dots) respondents in possible reasons for not engaging with USpeak. Means and 95% confidence intervals, N= 1459)

and Nerman 2014), we find that, at baseline, 6.6% of women responded to the poll compared to 12.3% of men. Thus in the baseline setting in which ICT is used to communicate with elites one’s preferences, but invitations are weak and impersonal, we find evidence of a large and significant gender gap in participation, similar to the one observed in the uSpeak program.

Moving to the experimental setting, we begin with a simple cross-tabulation of the raw data by treatment and gender for our verified users subsample. We find that in the control condition, the response rate of registered

Table 5: Mechanism Experiment

	<i>Dependent variable:</i>		
		Provided policy input	
	Base	Primary	Secondary
	(1)	(2)	(3)
Invitation	0.021* (0.011)	0.021* (0.011)	0.019** (0.008)
Flattening (Male*Invitation)		0.0004 (0.024)	-0.003 (0.017)
Observations	2,717	2,717	3,957
R <sup>2</sup>	0.153	0.153	0.165
Adjusted R <sup>2</sup>	0.108	0.107	0.115
<i>Note:</i>	<i>Male normalized to have 0 mean. *p &lt; 0.1</i>		

male users was almost double that of female users (11.2% as against 5.7%). Moving to the personalized invitation treatment, for both male and female UBridge registered users, overall response rate has increased by about 2 percentage points (13.4% against 7.5%). In order to more formally test whether there is a differential effect of the invitational treatment by gender we implement analyses pre-registered at EGAP’s registration web page.<sup>23</sup> Results, reported above in Table 5 (columns 2 and 3), confirms that the invitation effect was almost identical across men and women.

We conclude that though personal invitations can have a powerful effect on rates of participation, there is no evidence here that they have a differential effect by gender. It remains to be explored in future work whether invitational effects have a differential effect for other forms of marginalization.

## 6 Conclusion

This study integrates three related field experiments designed to assess whether innovations in information communication technologies (ICTs) can be harnessed to improve weak political communication, prevalent in many low-income countries. Evidence from a framed field experiment (FFE) conducted before rolling out a national program suggested not only that there is underlying demand to contact representatives using mobile technology, but also that ICTs have a genuine potential to increase levels of political engagement in a way that flattens access for marginalized populations. By contrast, when brought to scale using a natural field experiment (NFE) implemented nationwide, we find significantly lower levels of citizen engagement, with marginalized populations especially refraining from using the ICT platform to raise voice. These results have implications for theory, policy, and research methodology.

Our study contributes primarily to our understanding of the promises and pitfalls of ICT-based political communications, at least in the context of low-income countries. Consider four findings that help account for the weak uptake of Uganda’s national parliamentary communication system.

<sup>23</sup>Note the primary and secondary analyses are as well registered. The base column (1) is added to provide information on the unconditional invitation effect; i.e., without including heterogeneous gender effects.

First, we learned from the FFE that a nontrivial share of citizens, including especially marginalized citizens, want to communicate with their representatives in government using new technological innovations, and are willing to pay to do so. This stands in contrast to accounts of disengagement as reflecting alienation or apathy. We also know that many—though clearly not all (see Figure 5)—have the capacity and means to do so. The results from the FFE support the idea that mobile technology could, under the right conditions, change the relationships between voters and representatives in the developing world. An examination of the scaled up system alone would have masked this core insight.

Second, from an experimental manipulation in the NFE we found little evidence that the differences across field experiments is due simply to scale effects. Specifically, we do not find evidence that system usage is a strategic response to how many others are contacting their MP via the ICT platform. We believe that improving our understanding of the conditions under which constituents might view IT-based communication with public officials as complements or substitutes, is an important research avenue for future work to explore.

Third, complier analysis suggests that although the “first stage” of treatment using radio ads was weak, given the scale of the experiment it was far from being weak enough to account for the low engagement. Indeed, the estimated complier effect in the NFE is about 4% of the effect observed in the FFE.

Fourth, from the ‘mechanism’ experiment we learned that there is a reasonably strong responsiveness to personal invitations to engage politically when interest articulation is at stake, but we do not find evidence of the kind of differential responsiveness that would be needed to account for differences in flattening effects across experimental settings.

These findings suggest that the disappointing results of the uSpeak program are *not driven by weak demand*. In contrast, survey evidence suggests weaknesses in the marketing of the system itself. While a relatively large number of constituents were exposed to the radio ads, citizens had difficulty retaining and internalizing the information needed for acting conditional on hearing about the service. Moreover, our analysis suggests that *agent effects*—i.e., that the change in the identity of the implementer, which was easily observed by experimental subjects—likely have been very consequential. Specifically, we find strong evidence that general trust in the responsiveness of politicians is preventing engagement but is also rational. Interestingly in our case, agent effects do not stem from motivation differences between implementers (as for example identified by Berge et al. (2012)), but rather from the way agent identities interact with citizen expectations. In Uganda, as in many electoral authoritarian regimes—the most common regime type in Africa—low levels of political efficacy are discouraging political action; ICT innovations, by themselves, cannot force non-responsive politicians to become responsive.

With the multiple pieces of evidence available to us we infer that the failure of the nationwide program is not simply a function of weak demand on the part of citizens or to the weakness of marketing mechanisms but is a function of larger inequalities. Some of these, such as unevenness in receipt of invitations from Parliament, might be addressable through improved interventions. However, some reflect more fundamental weaknesses in the broader political system, most notably cynicism regarding the competence and motivations of politicians, which Parliament likely cannot address easily through technological innovation.

Our study also has broader implications for research methodology, and especially for the extent to which outcomes of scaled-up programs can be gleaned from results of controlled small-scale interventions. The literature on scaling up has largely focused on assessing the extent to which experimental estimates in one context apply in another. Some of this literature highlights the problems in using a small handful of

studies as the basis for inferences to different contexts (Collaboration 2015). Other work highlights the costs of extrapolation. Comparing non-experimental and experimental estimates that rely on the same data, Pritchett and Sandefur (2013) conclude that non-experimental estimates with the same subject population can better predict treatment effects as compared to experimental results from other contexts, because contextual variation can drive bigger differences in the estimated effectiveness of a program than selection bias.<sup>24</sup>

Importantly, there should not have been great differences in the subject population between the FFE and the NFE. The FFE was offered to a random sample of subjects from every constituency in Uganda, while the scaled-up program was offered to a random sub-sample of 186 constituencies, out of a total of 238 (see Supplementary Materials, Figure 3).

Thus the differences we observe draw attention to a distinct problem, largely overlooked by the extant external validity literature: the external validity across the nuts and bolts of *interventions* and not necessarily across *populations*.<sup>25</sup> This kind of validity problem is especially critical when lessons from carefully controlled small-scale studies are intended to inform policies to be implemented at a larger scale. Our results provide a cautionary tale for researchers and policy makers seeking to make such claims.

In our analysis, we identified several distinct reasons why outcomes of experiments may fail to replicate when brought to scale. These include already well-appreciated effects that relate directly to scale (see also Deaton (2010) on general equilibrium effects). In addition, we highlight possible effects related to the changing agents involved when interventions are implemented at scale (see also Bold et al. (forthcoming) on capacity and motivation of implementing organizations), and we identify differences related to details in the design between controlled interventions and interventions implemented in the political wild, of the form that may be relevant for other studies.

Ironically when design details matter, a first response is to resort to controlled conditions to get those details right. This might be an appropriate approach when seeking to control for all factors but a manipulated variable of interest, but one core lesson from our study is that the importance of those details may only become apparent once researchers' control is removed.

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<sup>24</sup>There is a growing literature debating the tradeoffs associated with different approaches to generating out-of-sample predictions based on experimental data. Hotz, Imbens, and Mortimer (2005) suggest using subject's observed characteristics as predictive of treatment effects independent of context. Gechter (2015) proposes a method that uses differences in outcome distributions for individuals with the same characteristics and treatment status in the original study and the context of interest to learn about unobserved differences across context.

<sup>25</sup>Some of our divergent findings do relate to endogenous changes in populations as a consequence of the factors outside the control of the research team. For example, one reason we do not find price effects in the scaled-up program similar to those found in the controlled experiment, can be attributed to the fact that the national intervention was taken up by relatively well off and engaged citizens who are unlikely to be sensitive to a small price subsidy.

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